

RESEARCH ARTICLE

Towards Carbon Neutrality: The Impact of Private AI Investment and Financial Development in the United States – An Empirical Study Using the STIRPAT Model

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Abstract

This investigation analyses the influence of private AI investment and financial development (FD) on CO₂ emissions in the United States, using the STIRPAT framework to account for the functions of GDP, population, and foreign direct investment (FDI). The data's robustness was verified through the application of a variety of unit root tests, which confirmed that the variables are free of unit root issues and exhibit a varied order of integration. The ARDL bound test was used to investigate the cointegration among the variables and it found a long-run equilibrium relationship. The ARDL model results show that income, FDI, FD, and population significantly increase CO₂ emissions in both the short and long term. In contrast, we found that private investment in AI led to a significant reduction in CO₂ emissions over these time frames. Additional estimations were conducted using FMOLS, DOLS, and CCR methods to verify the ARDL results, all of which attested to the initial findings' robustness. In addition, the study implemented a pairwise Granger causality test to illustrate the directional relationships between the variables. There is a unidirectional causal link between GDP, private AI investment, FDI, population, and CO₂ emissions, according to the findings. Most notably, we observed bidirectional causality between CO₂ emissions and FD. Diagnostic tests further corroborated the validity of the study's conclusions, confirming that the model is free from specification errors, serial correlation, and heteroscedasticity.

Keywords: Private AI Investment; Financial Development; FDI; Carbon Neutrality; United States

Introduction

Researchers from all around the world have been discussing and analyzing environmental economics and global warming for a few decades (Abid et al., 2022). Growing urbanization and industrialization have brought about significant environmental changes in recent decades (Mehmood and Tariq 2020). People widely recognize CO₂ emissions as the primary proxy of ecosystem damage (Voumik et al., 2024). UN research predicts that the 2015 Paris Agreement's aims for limiting global disasters may not be achieved if emissions of CO₂ aren't drastically cut below 2°C and 1.5°C. Serious social and economic repercussions will follow the failure to meet the carbon mitigation target (Ridwan et al., 2024). With 6677 million metric tons of CO₂, CO₂ accounted for 81.3% of all GHG pollution in 2018 (EPA, 2020). Of the GHGs released in 2018, 22% came from industrial activity alone or predominantly (EIA, 2020). In 2021, the US possessed the largest economy and ranked second in terms of CO₂ emissions among developed nations (British Petroleum 2022; World Bank 2022). In this sense, AI approaches can predict CO₂ emissions, and contemporary technology can guarantee environmental sustainability. Moreover, Wang et al. (2019) demonstrated that there is no longer an association within CO₂ pollution and economic growth, as the US GDP increased by 19% between 2007 and 2016. Given the United States' large scale and high CO₂ emissions, further investigation is necessary to explore the correlation between CO₂ emissions and GDP (Ulussever et al., 2023). Therefore, it is required to demonstrate the relationship across private investment in AI (PAI), financial development (FD), and CO₂ emissions within the framework of the US, since this can potentially provide guidance for other countries. Therefore, we give special consideration to FD, FDI, and PAI due to their potential to increase CO₂ emissions in this country. Scholars have conducted a thorough analysis of the United States' ecological state of circumstances (Acheampong, 2018). Since it draws attention to the harmful effects that our behavior has on the ecosystem, we applied it as a demand-side measure of environmental degradation. The United States has set a goal to achieve carbon neutrality, which means reaching a state where the level of GHGs generated is equivalent to the amount eliminated by the atmosphere by no later than 2050. US President Biden announced the US's return to the Paris Agreement in 2021 with the Clean Energy Revolution and Environmental Justice Plan, which seek for a green energy economy and a net zero CO₂ emission by goal by 2050 (Harris, 2020). The state governments are eager to take some concrete steps against global warming to meet this objective (Dehdar et al., 2023). To achieve the SDGs, an economy must employ an effective policy to minimize CO₂ emissions in addition to controlling finances (Yu et al., 2024). Although the US boasts the eighth-highest GDP per capita among 189 countries, its financial system face significant challenges due to the overuse of natural assets and CO₂ emissions (Zafar et al., 2019). The expansion of GDP correlates with an upsurge in CO₂ pollutions (Raihan et al., 2024b). Moreover, industrial revolution led to a substantial rise in the use of AI, and its implication on energy usage and CO₂ releases could aid in the attainment of carbon neutrality. Furthermore, AI algorithms that decrease CO₂ emissions contribute to the emergence of the sustainable financial system. By implementing strategies such as optimizing industrial structure, updating networks, and supporting creativity in sustainable technologies can effectively cut CO₂ emissions. However, by enabling technology transfer to recipient nations, FDI promotes the formation of human resources and the spread of innovation (Shahbaz et al., 2019). Furthermore, it promotes the use of manufacturing goods that pollute the natural environment and spread industrial activity while also aiding in the GDP development of the host region (Balli et al., 2023). Furthermore, FD can encourage the purchase of expensive items like automobiles, houses, and other objects, leading to higher power use, manufacturing, and CO₂ emissions (Dogan and Turkekul, 2016). Conversely, when individuals rapidly depend on fossil fuels to enhance their quality of life, the planet suffers. Consequently, countries experiencing a surge in population will exhibit higher levels of CO₂ emissions (Esquivias et al., 2022). Our paper significantly contributes to the corpus of expertise currently available on ecosystem degradation in several important areas. First, it introduces a new method for analyzing short- and long-term implications on the US zone by breaking down the

STIRPAT model and combining it with the ARDL framework. We particularly scrutinize the possible determinants of CO₂ emissions, namely PAI, FDI, and FD. Furthermore, by employing a relatively new dataset that encompasses a thorough investigation from 1996 to 2022 in the United States, our work contributes to the abundance of research already available on the environmental effects of PAI, FDI, and FD. Second, we use traditional unit root tests like ADF, P-P, and DF-GLS to assess the factors' stationarity. Third, we implement the ARDL bounds technique to cointegration to determine if there exists a cointegrating connection between CO₂ emissions and their driving factors. We then use Granger causality tests to look at the causal connection among the selected factors. Lastly, we applied various approaches such as FMOLS, DOLS, and CCR, along with additional diagnostic instruments, to evaluate the validity of our results and address any potential issues with the dataset. This research shows that while GDP, FDI, and population all influence emission levels in the USA, FDI and PAI enhance environmental level by cutting CO₂ emissions. This document arranges the remaining sections in the following order: Part 2 briefly represents the prior literature, while the next subsection illustrates the empirical model and information. The 4th portion provides the methodology. The 5th part displays the findings and critiques, the 6th chapter provides the conclusion, and Section 7 finally discusses the policy implications.

Literature Review

Several current studies focus on the complex links between CO₂ emissions, GDP progress, FDI, AI investment, FDI, and population expansion in different locations globally. We review the current state of academic research in that area to identify any gaps in the existing body of knowledge. Moreover, we are trying to ascertain the novel and significant sides of our research that add to this continuously expanding field of study. Given the magnitude of global warming, it is essential to evaluate CO₂ emissions by employing a diverse array of mathematical models and theories (Ridwan et al., 2023). The consumption of non-renewable energy and the loss of forests are rising, which leads to climate change, there has been an enormous rise in the number of studies focusing on CO₂ emissions in recent years (Jaafar et al., 2020). Raihan et al. (2024c) conducted an experiment in Vietnam to look at the correlation between economic development and CO₂ release. They concluded that there is a direct connection between GDP growth and CO₂ emissions, using the DOLS technique. Raihan et al. (2022b) have established a clear correlation, indicating that in the United States, a 1% expansion in GDP causes a corresponding surge of 0.59% in CO₂ pollutions in the short term and 0.29% in the long run. Ahmad et al. (2024) utilized the ARDL bound test for cointegration to explore the impact of GDP on China's ecological degradation. They demonstrated that a 1% boost in GDP corresponds to a 0.51% increase in CO₂ emissions. In addition, Pattak et al. (2023) provide evidence that corroborates the findings of previous studies. They observe that a 1% expansion in Italian GDP over an extended time can leads to an 8.08% spike in CO₂ emissions. Raihan et al. (2023a) investigate the ecological impacts of China's nuclear energy usage from 1993 and 2022, focusing on the EKC and PHH. They demonstrated that the expectation is that reducing CO₂ emissions will improve environmental quality and accelerate GDP growth. Similarly, Mehmood et al. (2021) determined that the autonomous influence of GDP significantly reduces the CO₂ emissions of Bangladesh and Pakistan. Contrarily, Acheampong et al. (2022) employed the NARDL approach and found that there is a minimal distinction between the increase and decrease in GDP in relation with CO₂ pollutions. Numerous papers discussed the deployment of AI and ML to lessen CO₂ emissions in construction sectors (Peng, 2019). A sizable number of studies have employed patent filings as a stand-in for invention (Balsalobre-Lorente et al., 2018; Herrerias et al., 2016). AI models, such as ML models, use data to create strategies for reducing pollution of CO₂ from human activities (Delanoe et al., 2023). Similarly, Dong et al. (2023) adopted the dynamic panel data from China to create econometric models in order to look at the consequence of AI on CO₂ production. Their outcome demonstrates that AI greatly lowers CO₂ emissions. Conversely, empirical studies have demonstrated that robotic shocks significantly reduce CO₂ emissions in

China's manufacturing industry (Chen et al., 2022; Jiang et al., 2022). From 2005 to 2016, Liu et al. (2022) looked into how artificial intelligence affected China's carbon intensity. They illustrated that AI significantly lowers carbon intensity by applying the STIRPAT framework. Zhao et al. (2023) used a fixed effects model to investigate the mechanisms and consequences of AI on China's CO₂ pollutions between 2006 and 2019. The results indicate that there is a substantial drop of 6.63% in emissions for every 10% rise in the use of AI, suggesting that AI has the potential to considerably lower the intensity of pollutant emissions.

The hyperlink between FD and the quality of ecosystems has drawn the focus of a greater variety of scholars in recent years (Acheampong, 2019). The theory of FD emphasizes the impact of the monetary system on GDP growth and the potential of financial assets to contribute to sustainable development (Aghion et al., 2005). Mehmood (2024) uses the CS-ARDL methodology to evaluate the consequences of FD on CO₂ emissions in Pakistan, India, Bangladesh, and Sri Lanka over the period 1984–2017. They assert that FD is essential for South Asian countries' development in order to achieve carbon neutrality. Tamazian and Rao (2010) found that FD has a major role in determining environmental performance, based on their analysis of 24 transition economies. Zafar et al. (2019) conducted a study where they examined the factors affecting the state of the earth in OECD zones. They identified a negative association between FD (which stands for a specific variable) and CO₂ emissions. This means that as FD increases, the ecological health of the environment would deteriorate due to increased pollution. Nevertheless, Rjoub et al. (2021) noted that FD made more environmental pollution of Turkey. Suhrab et al. (2023) analyze the impact of FD on CO₂ emissions in Pakistan using annual time series data from 1985 to 2018. The conclusion demonstrated a positive correlation between heightened FD and elevated levels of CO₂ pollutions. At the same way, Shahbaz et al. (2023) examine the effects of FD and GDP growth on ecological health in 10 economies with the greatest natural impacts. The findings suggest that FD has a harmful effect on an ecosystem. Several studies, such as Shoaib et al. (2020) in D8 and G8 and Kihombo et al. (2021) in West Asia and the Middle East, have identified a direct correlation between financial progress and environmental harm. FDI is acknowledged as an essential environmental variable that boosts efficiency and GDP by progressing innovation and capital creation (Alvarado et al. 2017). The influence of FDI on ecological health varies across countries and regions, leading to conflicting findings from different research. FDI inflow, in particular, has been associated with increased ecological sustainability, primarily through technological knowledge and spillover channels (Duan and Jiang, 2021). It reduces ecosystem damage by supporting creative uses of green technologies, according to data from numerous studies (Zhu et al., 2018; Udemba et al., 2020; Lahiani, 2020). Using the ARDL method, Saadaoui et al. (2024) evaluate how FDI affected Turkey's CO₂ emissions from 1985 to 2021. In the long run, the conclusions indicate that FDI mitigates emission level of CO₂. Comparably, Zhang and Zhou (2016) demonstrated that FDI helps China reduce its CO₂ pollutions. Wei et al. (2019) assert that FDI has the potential to enhance the natural environment by reducing pollution, but the use of green total productivity factors cannot achieve this boost. Nonetheless, increases in financial investment might encourage industrialization, which could worsen the environment and increase pollution (Baloch et al. 2019). Salahuddin et al. (2018), on the other hand, studied the empirical effects of FDI on CO₂ emissions in Kuwait covering 1980 and 2013. Using the ARDL bounds testing methodology, they discovered that FDI increases CO₂ emissions in both short and long terms. In a similar vein, Sapkota and Bastola (2017) discovered proof of FDI's detrimental effects on the ecosystem and indicated that a 1% spike in FDI is responsible for a 0.04% increase in ecosystem damage in Latin America. Population expansion (POP) has become a serious issue for environmental pollution, and several academicians are focusing on the nexus between boost in population and the release of CO₂. It is one of the key elements causing pollution (Grigg, 1991). By deploying the sophisticated ARDL and the STIRPAT model, Voumik et al. (2023a) calculated the effects of rise in POP on CO₂ pollution. The empirical data show that growing populations can raise the nation's emissions of CO₂ in Kenya. According to Pachiyappan et al. (2021), there will be a 1.4%

rise in CO2 emissions in India. Furthermore, Voumik and Ridwan (2023) conducted an investigation in Argentina from 1972 to 2021 using the ARDL methodology, and the empirical conclusions revealed that population development degrades ecosystem level in the long run. In a similar way, several studies, such as Ali et al. (2020) in Malaysia and Zhang et al. (2023) in the top 10 nuclear-generating economies, found the same outcomes. But Alam et al. (2016) investigated how growth in populations in China, Brazil, India, Indonesia, and China affected CO2 emissions between 1970 and 2012. They demonstrated a statistically significant correlation between CO2 emissions and POP increase for Brazil and India, but a statistically insignificant one for China and Indonesia. Conversely, Begum et al. (2015) surprisingly showed that the pace of population growth in Malaysia had no apparent consequences on the country's per capita emissions of CO2. In contrast, Sulaiman et al. (2018) employed the ARDL model and found that population growth could only have an immediate implication on CO2 emissions in Nigeria. With its focus on the U.S. setting and its distinct socioeconomic and ecological features, this study seeks to close a large knowledge gap. Despite growing global attention to equitable development, few thorough studies in USA to look at the combined influences of population rise, FD, FDI, PAI, and GDP development on CO2 emissions. In-depth assessments of the intricate interactions between these factors tend to be absent in the literature that is currently available, especially when it comes to the STIRPAT model. The outcomes of the connection among FD, PAI, and CO2 emissions are still contradictory, even though earlier research acknowledged the necessity for more thorough investigations on these associations. But it's crucial to consider that financial innovation and private investment in AI can encourage the uptake of cutting-edge, environmentally friendly technologies, which could assist in a decline in CO2 pollutions. Therefore, the objective of this study is to fill these gaps and provide policymakers with useful information so they can establish long-term plans to cut emissions.

Methodology

Data and Variables

Table 1 displays a full description of the data sources and relevant information. This research uses a balanced time series dataset for the United States, spanning the years 1996–2021. The Global Footprint Network (GFN) provides statistics for assessing ecological viability through CO2 emissions, which act as the dependent variable. Our World in Data provides one important parameter in this analysis: private investment in AI. The IMF provides statistics for the FD which is another important factor. The World Development Indicators (WDI) also provides population, GDP, and foreign direct investment data. We select the information sources based on their accessibility.

Theoretical Framework

The research utilizes modified IPAT and STIRPAT models for analyzing data. Ehrlich and Holdren (1971) first suggested the IPAT model as a framework for examining how growing population harms the environment. They adopted the following set of model instances:

$$I \equiv P.A.T \dots\dots\dots (1)$$

Here, "P" stands for population size, "A" for affluence, "T" for technological level, and "I" for the impact on the environment. There are several issues with the original IPAT identification. Dietz and Rosa (1994, 1997) suggested a modified format regarded as STIRPAT to address these issues. It includes a stochastic equation to

account for unintentional errors in parameter estimations. The model not only incorporates factors beyond the original IPAT framework, but also enables the estimation of these factors' elasticities (Shu et al., 2024). These researchers examine the following formulation:

$$I_{it} = CP_{it}^{\gamma_1} A_{it}^{\gamma_2} T_{it}^{\gamma_3} \varepsilon_{it} \dots \dots \dots (2)$$

At time t, P represents the country's population, A represents its wealth, and T represents its technology. The constant term in the STIRPAT model is C, while the random error component is represented by ε. Conversely, the symbols γ₁, γ₂, and γ₃ denote the coefficients of P, A, and T, respectively. We can express the logarithmic transformation of the model as follows:

$$LnI_{it} = C + \gamma_1 LnP_{it} + \gamma_2 LnA_{it} + \gamma_3 LnT_{it} + \varepsilon_{it} \dots \dots \dots (3)$$

The variables used in our research were waste recycling as an expression of impact (I), population growth as an indicator of population (P), affluence (estimated by GDP, FD, and FDI), and technological level (as private investment in AI). The equation can be written as:

$$CO_2 = f(GDP_{it}, PAI_{it}, FD_{it}, FDI_{it}, POP_{it}) \dots \dots \dots (4)$$

Table 1. Source and Description of Variables

Variables	Description	Logarithmic Form	Unit of Measurement	Source
CO ₂	CO ₂ Emission	LCO ₂	CO ₂ Emission (kt)	GFN
GDP	Gross Domestic Product	LGDP	GDP per capita (current US\$)	WDI
PAI	Private Investment in Artificial Intelligence	LPAI	Estimated Investment in AI (US\$)	Our World in Data
FD	Financial Development	LFD	Financial Development Index	IMF
FDI	Foreign Direct Investment	LFDI	Net Inflows (Current US\$)	WDI
POP	Population	LPOP	Population, total	WDI

The explanatory variables in this instance are GDP, PAI, FD, FDI, and POP, whereas the dependent variable is CO₂. An alternate way to describe the empirical model in logarithmic form is as follows:

$$LnCO_{2it} = \beta_0 + \beta_1 LnGDP_{it} + \beta_1 LnPAI_{it} + \beta_2 LnFD_{it} + \beta_3 LnFDI_{it} + \beta_4 LnPOP_{it} + \varepsilon_{it} \dots \dots \dots (5)$$

Where, β₀ is the intercept term and β₁, β₂, β₃, and β₄ are the coefficients of selected independent variables accordingly. The letter ε and L, stand for the natural log and the model's error term.

Empirical Methods

In this research, the DF-GLS, PP, and ADF assessments are among the most frequently used tests for unit root. Secondly, the statistical technique will employ robustness evaluations (FMOLS, DOLS, and CCR) to ensure the accuracy of the results, as well as the ARDL method to analyze cointegration across short and long periods. We then implemented the Granger causality test to explore the connection among the components, followed by a multitude of diagnostic tests.

Unit Root Test

Panel unit root testing is one approach to determining the integrational pattern for each variable, as the order incorporates all factors (Ozturk and Al-Mulali, 2015). It guarantees that the series is stationary and delivers an approximation of the regression equation employing stationary approaches (Raihan et al., 2023b). The standard unit root assessments that determine the sequence of variable integration include the P-P (Phillips and Perron 1988), the ADF (Dickey and Fuller 1979), and the DF-GLS test developed by Dickey-Fuller (Elliott et al., 1992). In comparison to the Dickey-Fuller (DF), the ADF technique is more robust and suited to more complicated procedures (Fuller, 2009).

ARDL Methodology

We use the cointegration and bounds testing approaches to determine whether there is a long-term link among CO2 emissions and their determinants, following the findings of the unit root analyses. Pesaran et al. (2001) recommended that the ARDL bounds test be deployed in this investigation to document the sequence's cointegration. Its main advantage is the ability to evaluate both short and long-term factors simultaneously. This allows it to analyze a wide range of time series data without requiring extensive previous testing (Raihan et al., 2022a). Furthermore, this framework (Raihan et al., 2024a) allows for the application of I (0) or I (1), or any frictionally integrated time series variable these variables typically represent. We represent the ARDL bound test in Eq. (6) as follows:

$$\begin{aligned} \Delta LCO_{2t} = & \alpha_0 + \beta_1 LCO_{2t-1} + \beta_2 LGDP_{t-1} + \beta_3 LPAI_{t-1} + \beta_4 LFD_{t-1} + \beta_5 LFDI_{t-1} + \beta_6 LPOP_{t-1} \\ & + \sum_{i=1}^q \delta_1 \Delta LCO_{2t-i} + \sum_{i=1}^q \delta_2 \Delta LGDP_{t-i} + \sum_{i=1}^q \delta_3 \Delta LPAI_{t-i} + \sum_{i=1}^q \delta_4 \Delta LFD_{t-i} + \sum_{i=1}^q \delta_5 \Delta LFDI_{t-i} \\ & + \sum_{i=1}^q \delta_6 \Delta LPOP_{t-i} + \varepsilon_t \end{aligned} \tag{6}$$

In the ARDL limit analysis, the F-distribution and critical values proposed by Pesaran and Timmermann (2005) are used within the framework of the equation denoted by (6), where q represents the optimal lag length and Δ denotes the first difference operator. Equation (6) serves as the first step in the estimating process. Pesaran et al. (2001) suggest that long-term correlations with F-statistics that fall between the threshold values are inconclusive, but those that fall below the threshold should receive acceptance.

Robustness Check

Phillips and Hansen (1990) recommended the FMOLS examination, Park (1992) reported the CCR test, and Stock and Watson (1993) calculated the DOLS test to assess the robustness of the ARDL outcomes. Before implementing FMOLS and DOLS, the CCR (Canonical Cointegrating Regression) method serves as an excellent tool for verifying their accuracy (Sultana et al., 2023). The development of these methodologies addressed two primary concerns. Before implementing FMOLS, DOLS, or CCR, it is necessary to meet the cointegration criteria among the I(1) parameters (Raihan et al., 2023b). A second benefit of these methods is that they address issues such as serial correlation, bias from absent variables, heterogeneity, and measurement errors (Raihan and Tuspekova 2023).

Granger Causality Test

The Granger causality test is paired and evaluates the sum of the past and current values of the independent variable (X) and dependent variable (Y) (Voumik et al., 2023b). The same is true for Y and X's causal link; if the outcomes deviate from zero, then both parties are causally involved (Rahman and Majumder, 2022). The analysis used the paired Granger causality test introduced by Granger (1969), and the causal connection between X_t and Y_t is depicted in the following equation.

$$E(Y_{t+h}|J_t, X_t) = E(Y_{t+h}|J_t) \quad (9)$$

Here, J_t notation is used for the sets of information gathered from all of the outcomes up to a certain point of time (t).

Diagnostic Test

The Lagrange Multiplier (LM), Jarque-Bera (Jarque and Bera, 1987), and BPG (Breusch and Pagan, 1979) tests are important in time series analysis for making sure that model assumptions are correct and that results are stable. Since many statistical models require normally distributed errors for proper inference so the Jarque-Bera investigation can be performed. By looking for serial correlation in residuals, the LM examination ensures that errors do not converge over time and produce skewed and misleading estimations. Finally, we utilized the BPG test to confirm the heteroscedasticity, or nonconstant variance, of the residuals.

Results and Discussion

Table 2 summarizes information on a variety of factors. The average LCO2 is 15.46, with a range from 15.27 to 15.56. The mean LGDP is 10.64, with a minimum of 10.08 and a maximum of 11.15. Notably, LFDI has the highest mean (25.89), while LFD has the lowest. The standard deviations of LCO2 (0.08), LGDP (0.31), LPAI (1.66), LFD (0.13), LFDI (0.76), and POP (0.08) demonstrate the variability around the mean. LPAI is positively skewed, while LCO2, LGDP, LFD, LFDI, and POP are inversely skewed. All variables have a normal distribution with low Jarque-Bera probability values. The dataset includes 32 observations for each variable from 1990 to 2022 in the United States.

Table 2. Descriptive statistics of Variables

Variable	LCO2	LGDP	LPAI	LFD	LFDI	LPOP
Mean	15.46444	10.64393	22.0143	-0.167379	25.89841	19.4995
Median	15.45192	10.71885	21.2377	-0.096679	26.09007	19.50906
Maximum	15.56919	11.15938	25.66873	-0.081949	26.96048	19.62079
Minimum	15.27889	10.08116	20.55212	-0.520773	24.13474	19.33546
Std. Dev.	0.080244	0.318778	1.665446	0.137067	0.768762	0.086797
Skewness	-0.466593	-0.255693	0.989543	-1.662635	-0.716141	-0.311024
Kurtosis	2.635357	1.888894	2.466317	4.273768	2.616801	1.894884
Jarque-Bera	1.338404	1.994763	5.602132	16.90654	2.931032	2.144299
Probability	0.512117	0.368844	0.060745	0.000213	0.230959	0.342272
Sum	494.8622	340.6058	704.4575	-5.356143	828.749	623.9841
Sum Sq. Dev.	0.199614	3.150205	85.98507	0.582405	18.32085	0.233545
Observations	32	32	32	32	32	32

Three stationarity tests for log-transformed variables at the level and first difference form are shown in Table 3. It seems that only the population factor is stationary at level I(0) at the 1% significance thresholds in each of the three unit root evaluations. Before we took into account their initial differences, the CO2, GDP, foreign direct investment, private investment in AI, and FD were non-stationary and significant at 1% significance thresholds. Thus, the ARDL methodology should be used to conduct the evaluation today because of the heterogeneous sequence of integration.

Table 3. Results of Unit root test

Variables	ADF		P-P		DF-GLS		Decision
	I(0)	I(1)	I(0)	I(1)	I(0)	I(1)	
LCO ₂	-0.155	-4.954***	-0.371	-4.959***	-0.521	-3.980***	I(1)
LGDP	-0.878	-4.841***	-0.953	-4.829***	-0.881	-4.085***	I(1)
LPAI	-0.806	-7.505***	-1.897	-7.403***	-0.673	-6.194***	I(1)
LFD	-2.731	-4.381***	-2.087	-4.551***	-1.889	-4.650***	I(1)
LFDI	-1.793	-6.508***	-1.565	-6.615***	-0.715	-4.882***	I(1)
LPOP	-4.896***	-7.881***	-6.881***	-8.605***	-3.831***	-5.750***	I(0)

Table 4 presents the bound analysis outcomes and reports the F-statistic as 4.2150, serving as a test statistic. We categorize the critical values using the integration order of the variables (I (0) and I (1)). The critical values for I (0) and I (1) are 2.08 and 3, accordingly, at the 10% significance level. Similarly, we provide critical values for both integration orders at the 5%, 2.5%, and 1% significance levels. These demonstrate the existence of a long-term correlation across the selected indicators.

Table 4. Results of ARDL bound test

Test Statistic	Value	Signif.	I(0)	I(1)
F-statistic	4.2150	10%	2.08	3
k	5	5%	2.39	3.38
		2.50%	2.7	3.73
		1%	3.06	4.15

We can evaluate their long-term relationship once the bound testing procedure reveals their cointegration. Table 5 adopts the novel ARDL method to outline the effects of different variables on LCO2 in the US, both in the short and long run. We can see from the given information that, a 1% spike in economic growth, the ecological condition drops by 0.057% over the long and by 0.044% over the short term. Beside the LGDP coefficient is both positive and statistically significant, we can infer that environmental pressure rises as US economic expansion accelerates. Azam et al. (2022), Cheng et al. (2021), Zafar et al. (2022), and Chien et al. (2023) have corroborated the findings and validated the beneficial association between GDP growth and carbon pollution. Furthermore, Pradhan et al. (2024) noted that as the economy expands, so does the consumption for goods and services, which raises output and, consequently increases emissions. According to the estimated coefficient for LPAI, there is an upward trend between LCO2 and private investment in AI, which promotes ecological condition in the US. In particular, an increase of 1% in LPAI delivers a long-lasting cut in CO2 emissions of 0.056% and an immediate decrease of 0.027%. This is because inventions are critical for improving energy efficiency and lowering CO2 emissions. Khalid et al. (2023) and Shahbaz et al. (2020) supported our conclusions. Additionally, higher amounts spent on eco-friendly procedures and greener technologies lessen their destructive implications on ecosystem and improve the health of our planet (Yao et al., 2021).

Table 5. ARDL short-run and Long-run Estimation

Variable	Coefficient	Std. Error	t-Statistic	Prob.
Long-run Estimation				
LGDP	0.057	0.6525	0.0882	0.012
LPAI	-0.056	0.0220	-2.5419	0.023
LFD	-0.357	0.2288	-1.5627	0.003
LFDI	0.027	0.0364	2.7630	0.025
LPOP	0.310	0.6517	2.1170	0.031
C	10.477	5.1367	3.2321	0.000
Short-run Estimation				
D(LCO2(-1))	0.127	0.1120	1.1368	0.071
D(LGDP)	0.044	0.2061	5.0596	0.001
D(LPAI)	-0.027	0.0086	-3.1973	0.015
D(LFD)	0.214	0.1352	2.5888	0.034
D(LFDI)	0.007	0.0106	-4.6978	0.000
D(LPOP)	0.2983	1.4912	2.5412	0.012
CointEq(-1)*	-0.475	0.1038	-4.5801	0.003

The calculated coefficients for LFD show an obvious inverse correlation with LCO2. This demonstrates that a 1% rise in FD generated a 0.214% short-term spike and a 0.357% long-term reduction. For this, the financial success plays a significant role in maintaining environmental sustainability in US. Numerous studies have revealed that the growth of the financial industry enhances the natural health (Saud et al., 2018; Latif and Faridi, 2023; Kartal et al., 2023). However, Pata et al. (2023) found the opposite result, holding that FD in the US had no bearing on the quality of the ecosystem. In addition, the destructive and statistically significant indications of the FDI coefficients suggest that an expansion in LFDI in both the long and short-run has a detrimental effect on environmental quality. Specifically, a 1% expansion in LFDI will result in a 0.027% and 0.007% increase in LCO2. It concludes that the present foreign direct investment of the United States is not conducive to reducing pollution; moreover, FDI can uplift the pollution level through energy consumption (Yang et al. 2020). Our findings align with Boubacar et al. (2024) in Africa, Zhang et al. (2023) in China, and Kouassi et al. (2024) across 43 African nations; they also concluded that greater FDI causes more release of CO2. Similarly, the results presented in Table 5 reveal that increased population degrades environmental quality in the US. An additional 1% expansion of LPOP will cause a 0.310% long and 0.298% short-run rise in CO2 pollutions. Our result aligns with Rehman et al. (2022) in Pakistan, Anser et al. (2020) in SAARC countries, and, Rehman et al. (2023) from a global perspective.

The inferences from the ARDL investigation are backed by the robustness findings shown in Table 6. At 1% significance threshold, the LGDP coefficients in the FMOLS, DOLS, and CCR models are statistically significant. According to FMOLS and CCR estimations, a 1% development in GDP results in a 0.196% and 0.117% boost in carbon pollutions, respectively, while DOLS reports a 0.016% cut in carbon emissions. A 1% rise in LPAI generates a corresponding drop in LCO2 of 0.045%, 0.030%, and 0.047% in FMOLS, DOLS, and CCR. In contrast, this result is significant in DOLS at the 10% level, as well as in FMOLS and CCR at the 1% level. Furthermore, in the FMOLS model, a 1% increase in LFD resulted in a 0.490% drop in LCO2. Conversely, there prevails a positive correlation, with a rise in CCR of 0.484% and a spike in DOLS of 0.352%. Furthermore, using the three estimation approaches, a 1% increase in LFDI yields an upsurge in LCO2 of 0.012%, 0.013%, and 0.010%. On the other hand, in FMOLS, a 1% rise in LPOP implies a 0.976% plunge in LCO2, which is notable at the 5% level. By comparison, DOLS and CCR predict a noteworthy 10% level of growth in emissions of 0.023% and 0.644%, respectively. The ARDL findings are in line with these contradictory responses.

Table 6. Robustness Check

Variables	FMOLS	DOLS	CCR
LGDP	0.196***	-0.016***	0.117***
LPAI	-0.045***	-0.030*	-0.047***
LFD	-0.490***	0.353**	0.484***
LFDI	0.012***	0.013**	0.010***
LPOP	-0.976**	0.023**	0.644*
C	10.678***	8.679***	10.561***

Table 7 outlines the causal relationships between various variables. This analysis reveals that LGDP does not cause LCO2, as observed by an F-statistic of 4.9537 and a p-value of 0.0154. This result suggests that, at the 5% significance level, we can reject the null hypothesis of no causal link between LGDP and LCO2. Additionally, there is evidence of one-way causation from LPAI, LFD, and LPOP to LCO2, supported by p-values below the conventional significance level. Similarly, there is a one-way causality between LCO2 and LFD, indicating that changes in LCO2 can influence LFD. Consequently, we can reject the null hypothesis of no causal relationship

in these cases. On the other hand, p-values higher than the conventional limit of significance suggest that neither LLCO2 nor LGDP significantly contribute to LPAI, LFD, or LPOP, nor does LFDI contribute to LCO2. This implies that changes in LCO2 have no impact on LFD, population expansion, AI investment by the private sector, or economic growth. Furthermore, since LFFDI has no effect on LCO2, we cannot rule out the null hypothesis that these interactions lack a causal relationship.

Table 7. Granger Causality Test

Null Hypothesis	Obs	F-Statistic	Prob.
LGDP \neq LCO2	30	4.9537	0.0154
LCO2 \neq LGDP		0.8454	0.4413
LPAI \neq LCO2	30	4.7929	0.0173
LCO2 \neq LPAI		0.0127	0.9873
LFD \neq LCO2	30	5.5694	0.0123
LCO2 \neq LFD		0.2487	0.7817
LFDI \neq LCO2	30	2.9103	0.0731
LCO2 \neq LFDI		0.2741	0.0125
LPOP \neq LCO2	30	4.6044	0.0198
LCO2 \neq LPOP		0.6101	0.5512

Table 8 displays the results of three different diagnostic evaluations. Because the data are all conflicting, it is evident that no diagnostic method can completely rule out the null hypothesis. According to the Jarque-Bera test, the residuals exhibit a normal distribution with a p-value of 0.1074. Then, the Lagrange multiplier assessment shows that there is no serial correlation in the residuals, with a p-value of 0.4031 being higher than the usual level of significance. In the end, the BPG test confirms that there is no heteroscedasticity issue with the residuals, producing a p-value of 0.7459.

Table 8. The results of diagnostic tests

Diagnostic tests	Coefficient	p-value	Decision
Jarque-Bera test	2.3015	0.1074	Residuals are normally distributed
Lagrange Multiplier test	1.0361	0.4031	No serial correlation exists
Breusch-Pagan-Godfrey test	0.7459	0.3214	No heteroscedasticity exists

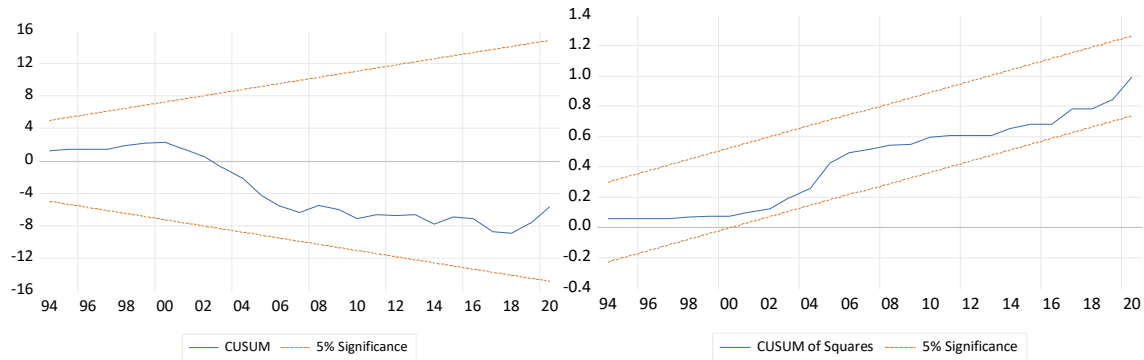


Figure 1. CUSUM and CUSUMSQ test

We performed the CUSUM and CUSUM-SQ tests to identify residuals of functions that reflect structural stability in both the short and long run. The statistical result in this case stays between the upper and lower bounds. The parameters are stable and well-defined at the 5% significant level, as evidenced by the CUSUMsq test plot's presence within the critical line (Figure 1).

Conclusion and Policy Recommendation

In the context of reaching carbon neutrality in the US, this study examines the connection among GDP growth, financial development, foreign direct investment, population growth, and private investment in AI. The analysis employs the ARDL bounds testing methodology within the STIRPAT framework using time series data spanning from 1996 to 2022. The unit root examinations confirm the stationarity of the variables and their lack of unit roots. Additionally, ARDL-bound experiments show that the factors do not cointegrate, confirming the lack of long-term equilibrium linkages. Nonetheless, the conclusions of the ARDL estimation indicate a strong and positive long-term correlation between GDP, FDI, POP, and CO2 emissions. The association between PAI and FD, on the other hand, is favorable and significant, suggesting a move toward more environmentally friendly financial practices and cutting-edge clean technology that lowers carbon emissions. According to the investigation, rising CO2 emissions in the United States are linked to economic development. However, technological and FD may be able to reduce these emissions, promoting healthy ecosystem. Robustness checks using FMOLS, DOLS, and CCR estimators validate the results, highlighting the complex relationships between these variables. Diagnostic tests confirm the regression model's reliability, and the residuals show no signs of serial correlation, heteroskedasticity, or deviations from normalcy. We also performed a Pairwise Granger causality test to investigate causal links between the variables. All things considered, this extensive investigation offers relevant details about the dynamics of development in finances, FDI, GDP expansion, AI investment, and CO2 emissions in the USA, providing a strong basis for responsible policy choices and environmental preservation plans. The findings have important policy ramifications for the United States as it strives to reconcile economic expansion with environmental sustainability. The relationship among wealth, FDI, FD, population, and CO2 emissions highlights the necessity for policies to tackle the environmental consequences of these economic activities. Law makers have to contemplate enforcing more stringent environmental restrictions on sectors that significantly contribute to CO2 emissions, particularly those that derive advantages from FDI and financial development. Moreover, the favorable influence of private investment in AI on decreasing CO2 emissions implies that advocating for AI-driven innovation might serve as a strategic means to alleviate environmental deterioration. One such approach is to provide incentives, such as tax exemptions, grants, or subsidies, to firms that invest in

AI technology with the goal of improving energy efficiency and decreasing carbon emissions. As a result, incorporating environmental risk assessments into financial decision-making processes and promoting green finance projects are critical measures to ensure that financial development is in line with sustainable environmental outcomes. Furthermore, given the one-way relationship in which GDP, FDI, and population influence CO₂ emissions, it is critical to develop economic strategies that separate economic expansion from environmental damage. This may entail advocating for the adoption of the circular economy, allocating resources towards renewable energy, and raising public consciousness regarding sustainable purchasing habits. Furthermore, it is critical to ensure that urban planning and population control strategies are in line with environmental sustainability objectives, especially in densely populated regions where the population increase is more likely to worsen CO₂.

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